# EE550 <br> Computational Biology 

Week 11 Course Notes

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## Topics

- Bioinformatics
- Preliminaries
- Randomness in measurements
- Probability distributions
- Histograms and empirical cumulative distributions
- Sample statistics
- Hypothesis testing using $t$ tests
- Parametric and nonparametric classification


## Motivation

- High throughput quantitative molecular biology data
- Cannot be processed or analyzed manually
- The data volume is well beyond the amount that can be handled manually
- Sequence data from many thousands of genes and proteins
- Signal transduction or gene transcription network maps
- Gene expression data from microarrays
- ...
- Manual analysis cannot provide any sense of statistical significance useful for making inferences regarding the biological problem at hand
$\rightarrow$ Computer algorithms


## Preliminaries

- Randomness in measurements
- All measurements are subject to fluctuations
- Fluctuations in the entity to be measured
- Transient effects
- Thermal noise in the measuring instrument
- Quantization errors
- Such fluctuations alter the measured value of a parameter of interest from its "true" value
- In other instances, the parameter of interest fluctuates in and of itself from one instance to another
- All these effects combine to produce deviations around some average


## Preliminaries

- Example: Cell-to-cell variation of the amount of CheR in E. coli chemotaxis
- When methylated, the receptor complex $X$ phosphorylates CheY that in turn triggers direction change
- The amount of CheR determining the steady state concentrations of the methylated receptor complex X changes from cell to cell
- As a result, some cells are more nervous and change direction more often, while others are much more relaxed
- All these effects combine to produce deviations around some average


## Preliminaries

## - Random variables

- Technically:
- A random variable is a mapping from a probability space $(S, \Omega, P)$ onto a measurable space ( $S, \Omega$ )
- $S$ is the domain; also called the universal set of all possible outcomes/values
- $\Omega$ is the sigma-algebra associated with the domain
- $P: \Omega \rightarrow[0,1]$ is the probability measure such that $P(S)=1$ and $P(\omega) \geq 0$ for all $\omega$ in $\Omega$
- Practically:
- A random variable denotes the values of a parameter of interest measured under noisy or erroneous but generally stable conditions
- The value of the random variable changes every time a measurement is made
- Ranges of possible values that a random variable can take are associated with a probability between 0 and 1


## Preliminaries

- Example:
- Consider a fair die
- A perfect cube with faces numbered from 1 to 6
- When thrown, it has equal chance to land on its different faces
- Throwing of this die corresponds to a random experiment
- The measurement related to this random experiment is the reading of the number written on the face looking up
- Each throw corresponds to a distinct realization of the random experiment
- The measurement is simply the outcome of the experiment
- Probabilities are assigned to collections (or sets) of events
- Q: Suppose a fair die is thrown. What are the chances that the outcome will be
- Greater than or equal to 1 ?
- Less than 10 ?
- Less than 100?
- 1 or 2 or 3 ?
- 4 or 5 or 6 ?
- 1 or 3 or 5 ?
- 2 or 4 or 6 ?
- 1 or 2 ?
- 2 or 4 ?
- 5 or 6 ?
- 1 ?
- 2?
- 3 ?
- ...


[^0]
## Preliminaries

- The odds of different possible outcomes are expressed in terms of probability distribution - mass or density - functions
- Let $X$ denote the random variable associated with the throwing of a fair die

$$
\begin{aligned}
& \operatorname{Pr}\{X=1\}=1 / 6 \\
& \operatorname{Pr}\{X=2\}=1 / 6 \\
& \operatorname{Pr}\{X=3\}=1 / 6 \\
& \operatorname{Pr}\{X=4\}=1 / 6 \\
& \operatorname{Pr}\{X=5\}=1 / 6 \\
& \operatorname{Pr}\{X=6\}=1 / 6
\end{aligned}
$$

- Therefore, the probability mass function of $X$, denoted by $p_{X}$, is

$$
p_{X}(x)=\left\{\begin{array}{cc}
1 / 6 & \text { if } x \in\{1,2,3,4,5,6\} \\
0 & \text { otherwise }
\end{array}\right.
$$



## Preliminaries

- The probability distribution of a random variable governs the odds of observing some specific values in a chance event
- In case the exact form of this probability is not known, it can be estimated
- using many realizations of the corresponding chance event
- A most common way of estimating underlying probability distributions is by way of histograms
- The more realizations, the better the estimate
- Still, ambiguities abound


## Preliminaries

- Consider estimating the underlying probability distribution of a fair die experiment from 100 independent realizations
- The die is thrown $N=100$ times
- The numbers that come up each time are recorded
- Let $N_{1}$ be the number of times the face with the number 1 comes up, and similarly for $N_{2}, N_{3}, N_{4}, N_{5}$, and $N_{6}$
- Or, simply, $N_{x}$ for $x=1, \ldots, 6$
- Clearly,

$$
N_{1}+N_{2}+N_{3}+N_{4}+N_{5}+N_{6}=100
$$

- Define $h$ by

$$
h(x)=\frac{N_{x}}{100}, x=1,2, \ldots, 6
$$

- Then, $h$ is a histogram of the 100 realizations of the random variable $X$, and an estimate of $p_{X}$


## Preliminaries

$N=100$

$N=100$




## Preliminaries

- The die throwing experiment describes a discrete random variable - The outcomes are elements in a finite set $\{1,2,3,4,5,6\}$
- More interesting examples tend to assume values from a continuum
- The random variables associated with such parameters are called continuous random variables
- Continuous random variables possess similar definitions as the discrete random variables
- Probability measures, chance events, ...
- But they differ in certain crucial ways, especially in how the probability distributions are defined
- Let $X$ denote the height of a freshman at IYTE in meters
- Q: What is the probability that a freshman at IYTE will be 1.70 m tall, i.e., $\operatorname{Pr}\{X=1.70\}=$ ?
- A: ZERO!!!
- But, but, but... A freshman does have a certain height; if it's not 1.70 EXACTLY, it is somewhere near...
- So what?


## Preliminaries

- The laws governing the chance structure associated with the values of continuous random variables are given in terms of set probabilities
- The probability of interest is not $\operatorname{Pr}\{X=1.70\}$, but $\operatorname{Pr}\{X \leq 1.70\}$
- The cumulative distribution function of $X$, denoted by $F_{X}(x)$, is defined as

$$
F_{X}(x)=\operatorname{Pr}\{X \leq 1.70\}
$$

- Note that

$$
\begin{aligned}
& -\lim _{x \rightarrow-\infty} F_{X}(x)=0 \\
& -\lim _{x \rightarrow \infty} F_{X}(x)=1
\end{aligned}
$$

- In turn, the probability density function $f_{X}(x)$ is defined as the derivative of $F_{X}(x)$ as

$$
f_{X}(x)=\frac{d F_{X}(x)}{d x}
$$

## Preliminaries

- There are certain key probability distribution families that have been found very useful in describing the chance structures associated with real life random events
- Gaussian probability distribution function
- Bell curve
- Exponential probability distribution function
- Time-to-event
- Binary probability distribution function
- Heads or tails?
- Binomial probability distribution function
- How many heads or tails in so many repeats?
- Poisson probability distribution function
- How many heads or tails so far?


## Preliminaries

- Gaussian probability distribution
- A continuous function with two parameters
- Mean $\mu$
- Variance $\sigma^{2}$
$f_{X}(x)=\frac{1}{\sqrt{2 \pi \sigma^{2}}} e^{\left(-\frac{(x-\mu)^{2}}{2 \sigma^{2}}\right)}$


## Preliminaries

- Exponential probability distribution
- Another continuous distribution, this time with one parameter
- The rate of change $\lambda$
- This is the only memoryless continuous distribution

$$
f_{X}(x)=\lambda e^{-\lambda x}
$$



## Preliminaries

- Binary probability distribution
- A discrete distribution with only two possible outcomes
$p_{X}($ "first outcome" $)=p$
$p_{X}($ "second outcome" $)=1-p$
- The set of outcomes can be varied
- $\{0,1\}$
- $\{-1,1\}$
- $\{A, B\}$
- ...



## Preliminaries

- Binomial probability distribution
- A discrete distribution counting two possible outcomes in so many independent repeats with

$$
\begin{aligned}
& p_{X}(\text { "first outcome" })=p \\
& p_{X}(\text { "second outcome" })=1-p
\end{aligned}
$$

- The probabilities are then given by
$\operatorname{Pr}\{" k$ first outcome in $n$ repeats" $\}$

$$
=\binom{n}{k} p^{k}(1-p)^{n-k}
$$





## Preliminaries

- Poisson probability distribution
- Another discrete distribution with one parameter
- Rate of change $\lambda$
- Counts the number of times an event of interest occurs in a fixed period of time

$$
p_{X}(x)=\frac{\lambda^{x} e^{-\lambda}}{x!}
$$

- Interestingly, the time separation between successive events is exponentially distributed



## Preliminaries

- A sample set represents a collection $\left\{x_{j}\right\}, i=1,2, \ldots, N$ of values observed from a given random variable
- A collection of freshman heights from a randomly selected group of 10 first year students
- Sequence lengths of 10000 randomly selected human proteins
- Ages (in years) of 120 Alzheimer's Disease patients
- The distribution of values in the sample set can be characterized using
- Histograms
- $N_{k}$ represents the number of samples in an interval $\left(x_{k-1}^{\prime}, x_{k}^{\prime}\right]$ with

$$
x_{0}^{\prime}<x_{1}^{\prime}<\cdots<x_{K-1}^{\prime}<x_{K}^{\prime}
$$

- $K$ represents the number of bins
- Sample statistics
- Sample mean $m=\frac{1}{N} \sum_{i} x_{i}$
- Sample variance $s^{2}=\frac{1}{N-1} \sum_{i}\left(x_{i}-m\right)^{2}$


## Preliminaries





Above: The probability density function of some random variable $X$

Right: Histograms of 1000 realizations of $X$ with different bin sizes


## Preliminaries

## - Remarks:

- Histograms are informative only when the bins are located and sized appropriately
- There is no sense in placing the bins on regions of zero occurrence
- If the bins are too small, the resolution will be high, but they will cover only a few samples producing large errors
- Larger bins will possess many samples providing a smaller error, but the resolution will be poor
- Sample mean $m$ and variance $s^{2}$ (standard deviation $s$ too) describe where the samples are centered and how wide they are dispersed
- This is usually fine for unimodal distributions with a single peak
- On the other hand, this is terribly inadequate to represent multimodal distributions
- The samples may be clustered around a handful of values with little or no dispersion
- The mean will not capture this localization, and the standard deviation will indicate large dispersion when there is only very little


## Hypothesis Testing

- Suppose we are given two sample sets

$$
\left\{x_{i}\right\}, i=1,2, \ldots, N_{x}
$$

and

$$
\left\{y_{j}\right\}, j=1,2, \ldots, N_{y}
$$

- The heights of freshman students in EE and MB\&G
- The sequence lengths of human and yeast proteins
- 
- The task is to decide if these two sample sets represent events with different characteristics
- These sample sets represent events with different characteristics if and only if the underlying probability distributions are different
- One option it to generate histograms for the two sets and see if they look different
- Feasible first-attempt, but difficult to infer a statistical significance measure
- Requires a measure of distance between histograms and permutation tests
- Another option is to assume these sample sets originate from distributions of the Gaussian family with potentially different parameters, and test to see if their parameters might be the same


## Hypothesis Testing

- Presumptions about the statistical nature of the observed data are tested against empirical evidence presented by the data
- Formally:
- A null hypothesis $H_{0}$ is formulated postulating a statement
- the uninteresting explanation for the observed data
- A complementary hypothesis $H_{c}$ is automatically formulated postulating the invalidity of the statement
- the interesting/desired/hoped-for explanation
- A probability $P$ is computed as the probability of observing the actual observed sample statistic under the null hypothesis
- If the probability is smaller than a prescribed significance threshold, the null hypothesis is rejected at the benefit of the complementary hypothesis
- Small $P$ values indicate that the sample statistic is unlikely to be observed if null hypothesis were true
- Typical $P$ value thresholds are $5 \%$ or $0.1 \%$
- Note that this strategy requires a statistic to be computed from the data with a known distribution under the null hypothesis
- Any statistic can be used as long as its distribution can be guessed well


## Hypothesis Testing Using a Two-Sample $t$-Test

- Consider the following problem:
- Two sample sets $\left\{x_{i}\right\}$ and $\left\{y_{j}\right\}$ are provided representing the value observed for a parameter of interest from two different groups
- $\left\{x_{i}\right\}$ are the realizations of a random variable $X$
- $\left\{y_{j}\right\}$ are the realizations of a random variable $Y$
- Let $\mu_{X}$ and $\mu_{Y}$ represent the unknown means of the random variables $X$ and $Y$
- The task is to test the validity of the null hypothesis

$$
H_{0}: \mu_{X}=\mu_{Y}
$$

with a significance threshold $\alpha \ll 1$
$\rightarrow$ two-sided two-sample $t$-test

## Hypothesis Testing Using a Two-Sample $t$-Test

- A $t$-test is a statistical comparison test that computes a probability for the null hypothesis given the data
- If the probability is smaller than the prescribed significance $\alpha$, the null hypothesis is rejected in favor of the complementary hypothesis
- A few variants exist for the $t$-test
- Equal sample sizes, equal variances
- Unequal sample sizes, equal variances
- Unequal sample sizes, unequal variances
- Paired vs. unpaired
- The test calculates a $T$ statistic for each case, and computes its probability when the null hypothesis is true as the $P$ value
- For unequal sample sizes, equal variances:

$$
T=\frac{m_{X}-m_{Y}}{s \sqrt{\frac{1}{N_{X}}+\frac{1}{N_{Y}}}}
$$

where

$$
\begin{aligned}
m_{X} & =\frac{1}{N_{X}} \sum_{i} x_{i}, m_{Y}=\frac{1}{N_{Y}} \sum_{j} y_{j} \\
s_{X}^{2} & =\frac{1}{N_{X}-1} \sum_{i}\left(x_{i}-m_{X}\right)^{2} \\
s_{Y}^{2} & =\frac{1}{N_{Y}-1} \sum_{j}\left(y_{j}-m_{Y}\right)^{2}
\end{aligned}
$$

$$
s=\sqrt{\frac{\left(N_{X}-1\right) s_{X}^{2}+\left(N_{Y}-1\right) s_{Y}^{2}}{N_{X}+N_{Y}-2}}
$$

and

$$
D F=N_{X}+N_{Y}-2
$$

## Hypothesis Testing Using a Two-Sample $t$-Test

- Procedure for testing for the equality of means:

1. Given the sample sets $\left\{x_{i}\right\}$ and $\left\{y_{j}\right\}$
2. Calculate the sample means and variances
3. Calculate the $T$ statistic
4. Compare the absolute value of the $T$ statistic to the critical value $T_{c}$ for which

$$
F_{t}\left(T_{c}\right)=1-\alpha / 2
$$

where $F_{t}$ denotes the cumulative distribution function of a $t$ random variable with the corresponding degrees of freedom under the null hypothesis

## OR

Calculate the $P$ value via

$$
P=2 \cdot\left(1-F_{t}(|T|)\right)
$$

and see if it is smaller than $\alpha$

## Hypothesis Testing Using a Two-Sample $t$-Test

- Equality of the means example:
- Let the sample sets be given as
$x$ 's: $\{0.83,-0.09,-0.46,0.05,-1.36,-0.21\}$
$y$ 's: $\{-0.08,0.13,1.94,0.57,0.27,0.99,0.41,0.87,0.35\}$
- Compute
- The sample means
$m_{x}=-0.2067$ and $m_{y}=0.6056$
- The sample variances
$s_{x}^{2}=0.5097$ and $s_{y}^{2}=0.3640$
- The $T$ statistic

$$
T=-2.3778
$$

- The $P$ value associated with this $T$ statistic is

$$
P=2 \cdot\left(1-P_{t}(|T|)\right)=2 \cdot 0.0167=0.0334
$$

- The critical value $T_{c}$ is

$$
T_{c}=2.1448
$$

- The null hypothesis is rejected!!



## Hypothesis Testing Using a Two-Sample $t$-Test

- Example (continued):
- Now, let the sample sets be given as $x$ 's: $\{0.83,-0.09,-0.46,0.05,-1.36,-0.21\}$
$y$ 's: $\{-0.08,0.13,(4.94,0.57,0.27,0.99,0.41,0.87,0.35\}$
- Compute
- The sample means
$m_{x}=-0.2067$ and $n_{y}=0.9389$;

- The sample variances
$s_{x}^{2}=0.5097$ and $s^{2}=2.3648$
- The $T$ statistic

$$
T=-1.6914
$$

- The $P$ value associated with this $T$ statistic is

$$
P=2\left(1-P_{t}(|T|)\right)=2 \cdot 0.0573=0.11
$$

- This time, the null hypothesis is not rejected!!
- What is going on??



## Hypothesis Testing Using a Two-Sample $t$-Test

- Remarks:
- The $t$ test is susceptible to deviations from the presumptions
- Gaussianity of the underlying distributions
- Presence of outliers
- In addition, it determines whether there is reason to believe that the unknown means are different, but says little about how different they are
- Given sufficient number of samples, the statistical power may suffice to detect even the tiniest differences between the means
- Conversely, not detecting a difference of the means in a significant manner may simply be because the available data does not provide sufficient statistical power to detect a small difference
- Finally, it is helpless when the random variables are multivariate
- Hotelling's $T^{2}$ test can be used but is problematic


## Parametric and Nonparametric Classification

- Often, several parameters are measured jointly and recorded in experiments
- Heights, ages, and grade point averages of college freshmen
- Lengths and amino acid compositions of amino acid sequences of human proteins
- Gene expression of 40 K genes in microarray experiments
- Such multivariate data sets require multivariate data analysis methods
- A common task when multivariate data from two or more sample sets are present is whether classification rules that distinguish these sets from one another can be constructed
- If such a rule can be constructed, one can then determine
- to which group a novel sample should belong
- which parameter values are critical to distinguish the samples of different groups and in what conditions
- Both these possibilities are absolutely vital to understand the biological problems in consideration


## Parametric and Nonparametric Classification

- Given training data, classifier construction strategies are studied under two general categories
- Parametric classification rules
- A parametric model is assumed for the underlying multivariate probability distributions of the different groups
- The parameters for these distribution models are estimated from available data
- An optimal decision boundary is deduced from the estimated probability distributions
- Nonparametric classification rules
- No parameter-based model is assumed
- Classification rules are constructed based on the similarity and distance structure between the available -manually annotated"training" samples


## Parametric and Nonparametric Classification

- Maximum likelihood classification
- Given the multivariate training data
- Estimate the means and the covariance matrices for all sample sets
- The estimated sample distributions then become multivariate Gaussian distributions with the corresponding mean vectors $\mu_{i}$ and the covariance matrices $\Sigma_{i}$ as

$$
f_{i}(\boldsymbol{x})=\frac{1}{\sqrt{(2 \pi)^{n}\left|\operatorname{det}\left(\Sigma_{i}\right)\right|}} e^{-\frac{1}{2}\left(x-\mu_{i}\right)^{T} \Sigma_{i}^{-1}\left(x-\mu_{i}\right)}
$$

- Construct the classification rule that assigns a new sample to the sample set with the greatest value of the probability density function at the new sample

$$
f^{\mathrm{ML}}(\boldsymbol{x})=\arg \max _{j} f_{j}(\boldsymbol{x})
$$

## Parametric and Nonparametric Classification



## Parametric and Nonparametric Classification

- Nearest neighbor classification:
- Store all the available data for training in a reference set $\left\{\boldsymbol{x}_{j}, y_{j}\right\}, \boldsymbol{x}_{j} \in I R^{n}$, $y_{j} \in\{1,2\}$ with $j=1,2, \ldots, \ell$
- Assign the newly observed sample to the class with most similar samples
- Similarity computed in terms of a defined measure, or as inverse distance, or a weighted combination, ...
- The classification rule is given by

$$
f^{\mathrm{NN}}(\boldsymbol{x})=y_{j_{0}}
$$

where $j_{0}=\arg \min _{j} \rho\left(\boldsymbol{x}, \boldsymbol{x}_{j}\right)$, with $\rho\left(\boldsymbol{x}, \boldsymbol{x}_{\boldsymbol{j}}\right)$ calculating the distance between samples $x$ and $x_{j}$

## Parametric and Nonparametric Classification



## Parametric and Nonparametric Classification

- Support vector machine classification:
- A maximum margin linear classifier is constructed to separate the samples of two different classes
- Nonlinear solutions are obtained by employing an inner product kernel to replace the original inner product between the samples
- polynomial, Radial Basis Function, sigmoid, ...
- Linear maximum-margin solution in the transform space corresponds to a nonlinear solution in the observation space
- For more details, see the literature
- Maximization of the margin using the method of Lagrange multipliers
- Karush-Kuhn-Tucker optimality conditions that produce the support vectors
- Generalization to multiple class problems


## Parametric and Nonparametric Classification



## Summary

- Bioinformatics uses statistical analysis techniques to address molecular biology questions emanating from quantitative data in large volumes
- The data collected from high throughput experiments can only be handled using computational methods
- These methods use different strategies to answer a variety of questions
- Whether the nature of measured parameters change from one group to another
- Whether it is possible to derive classification rules to distinguish the different groups based on the measured data


[^0]:    Source: https://www.123learning.co.uk/pack-of-10-dice

