# EE550 Computational Biology

Week 11 Course Notes

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# Topics

- Bioinformatics
  - Preliminaries
    - Randomness in measurements
    - Probability distributions
    - Histograms and empirical cumulative distributions
    - Sample statistics
  - Hypothesis testing using t tests
  - Parametric and nonparametric classification

# Motivation

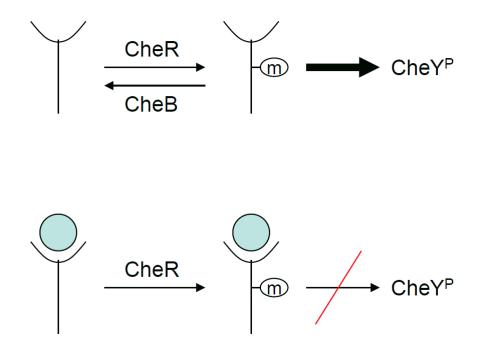
- High throughput quantitative molecular biology data
  - Cannot be processed or analyzed manually
    - The data volume is well beyond the amount that can be handled manually
      - Sequence data from many thousands of genes and proteins
      - Signal transduction or gene transcription network maps
      - Gene expression data from microarrays

- ...

- Manual analysis cannot provide any sense of statistical significance useful for making inferences regarding the biological problem at hand
- ➔ Computer algorithms

- Randomness in measurements
  - All measurements are subject to fluctuations
    - Fluctuations in the entity to be measured
    - Transient effects
    - Thermal noise in the measuring instrument
    - Quantization errors
  - Such fluctuations alter the measured value of a parameter of interest from its "true" value
  - In other instances, the parameter of interest fluctuates in and of itself from one instance to another
  - All these effects combine to produce deviations around some average

- Example: Cell-to-cell variation of the amount of CheR in E. coli chemotaxis
  - When methylated, the receptor complex X phosphorylates CheY that in turn triggers direction change
  - The amount of CheR determining the steady state concentrations of the methylated receptor complex X changes from cell to cell
  - As a result, some cells are more nervous and change direction more often, while others are much more relaxed
  - All these effects combine to produce deviations around some average



- Random variables
  - Technically:
    - A random variable is a mapping from a probability space (S, Ω, P) onto a measurable space (S, Ω)
      - *S* is the domain; also called the universal set of all possible outcomes/values
      - $\Omega$  is the sigma-algebra associated with the domain
      - $P: \Omega \to [0,1]$  is the probability measure such that P(S) = 1 and  $P(\omega) \ge 0$  for all  $\omega$  in  $\Omega$
  - Practically:
    - A random variable denotes the values of a parameter of interest measured under noisy or erroneous but generally stable conditions
      - The value of the random variable changes every time a measurement is made
      - Ranges of possible values that a random variable can take are associated with a probability between 0 and 1

- Example:
  - Consider a fair die
    - A perfect cube with faces numbered from 1 to 6
    - When thrown, it has equal chance to land on its different faces
  - Throwing of this die corresponds to a random experiment
  - The measurement related to this random experiment is the reading of the number written on the face looking up
  - Each throw corresponds to a distinct realization of the random experiment
    - The measurement is simply the outcome of the experiment
  - Probabilities are assigned to collections (or sets) of events

- Q: Suppose a fair die is thrown. What are the chances that the outcome will be
  - Greater than or equal to 1?
  - Less than 10?
  - Less than 100?
  - 1 or 2 or 3?
  - 4 or 5 or 6?
  - 1 or 3 or 5?
  - 2 or 4 or 6?
  - 1 or 2?
  - 2 or 4?
  - 5 or 6?
  - 1?
  - 2?
  - 3?
    - ...



Source: https://www.123learning.co.uk/pack-of-10-dice

- The odds of different possible outcomes are expressed in terms of probability distribution – mass or density – functions
  - Let X denote the random variable associated with the throwing of a fair die

$$Pr{X = 1} = 1/6$$
  

$$Pr{X = 2} = 1/6$$
  

$$Pr{X = 3} = 1/6$$
  

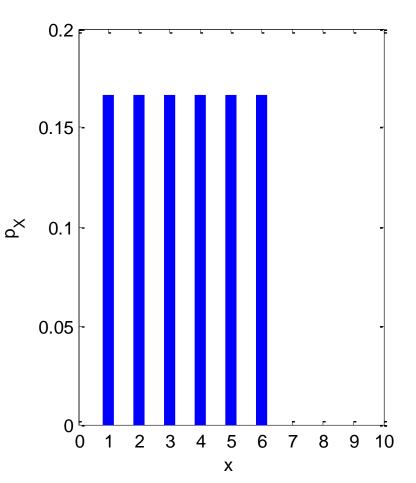
$$Pr{X = 4} = 1/6$$
  

$$Pr{X = 5} = 1/6$$
  

$$Pr{X = 6} = 1/6$$

- Therefore, the probability mass function of X, denoted by  $p_X$ , is

$$p_X(x) = \begin{cases} 1/6 & \text{if } x \in \{1,2,3,4,5,6\} \\ 0 & \text{otherwise} \end{cases}$$



- The probability distribution of a random variable governs the odds of observing some specific values in a chance event
- In case the exact form of this probability is not known, it can be estimated
  - using many realizations of the corresponding chance event
- A most common way of estimating underlying probability distributions is by way of histograms
  - The more realizations, the better the estimate
  - Still, ambiguities abound

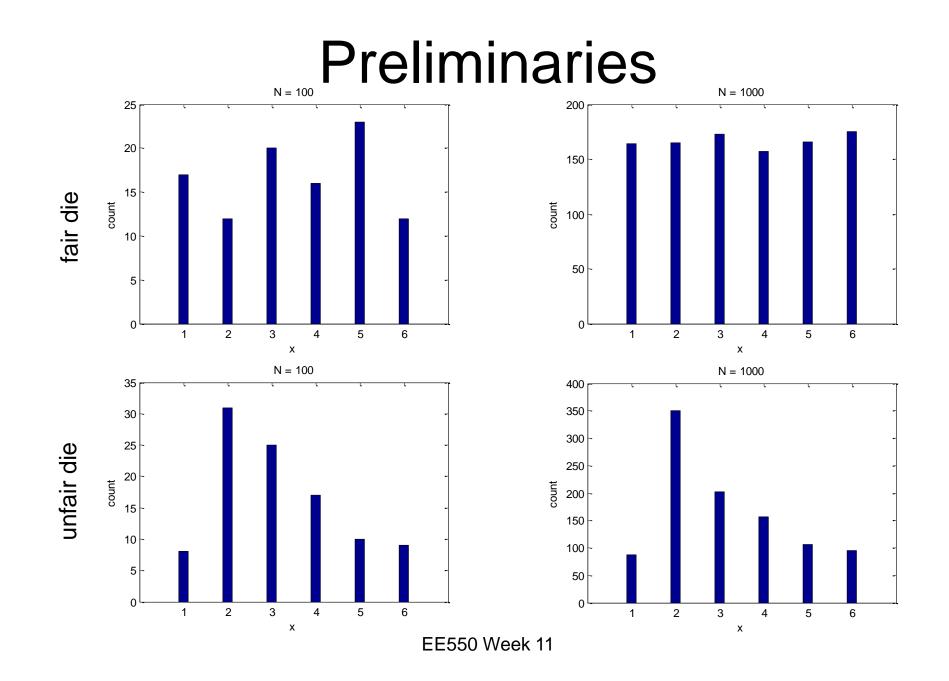
- Consider estimating the underlying probability distribution of a fair die experiment from 100 independent realizations
  - The die is thrown N = 100 times
  - The numbers that come up each time are recorded
  - Let  $N_1$  be the number of times the face with the number 1 comes up, and similarly for  $N_2$ ,  $N_3$ ,  $N_4$ ,  $N_5$ , and  $N_6$ 
    - Or, simply,  $N_x$  for x = 1, ..., 6
    - Clearly,

$$N_1 + N_2 + N_3 + N_4 + N_5 + N_6 = 100$$

– Define h by

$$h(x) = \frac{N_x}{100}, x = 1, 2, \dots, 6$$

– Then, *h* is a histogram of the 100 realizations of the random variable *X*, and an estimate of  $p_X$ 



- The die throwing experiment describes a **discrete** random variable
  - The outcomes are elements in a finite set {1,2,3,4,5,6}
- More interesting examples tend to assume values from a continuum
- The random variables associated with such parameters are called continuous random variables
  - Continuous random variables possess similar definitions as the discrete random variables
    - Probability measures, chance events, ...
  - But they differ in certain crucial ways, especially in how the probability distributions are defined
    - Let *X* denote the height of a freshman at IYTE in meters
    - Q: What is the probability that a freshman at IYTE will be 1.70m tall, i.e.,  $Pr{X = 1.70} =?$
    - A: ZERO!!!
    - But, but, but... A freshman does have a certain height; if it's not 1.70 EXACTLY, it is somewhere near...
    - So what?

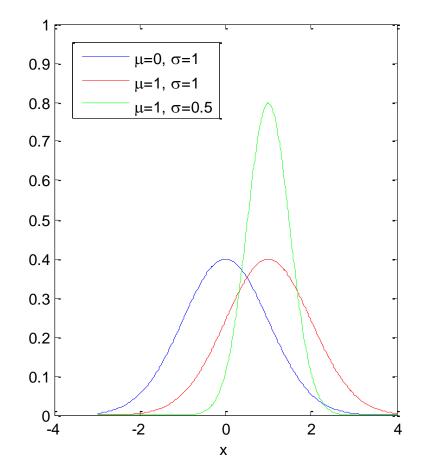
- The laws governing the chance structure associated with the values of continuous random variables are given in terms of set probabilities
   The probability of interest is not Pr{X = 1.70}, but Pr{X ≤ 1.70}
- The cumulative distribution function of *X*, denoted by  $F_X(x)$ , is defined as  $F_X(x) = \Pr\{X \le 1.70\}$
- Note that
  - $-\lim_{x\to-\infty}F_X(x)=0$
  - $-\lim_{x\to\infty}F_X(x)=1$
- In turn, the **probability density function**  $f_X(x)$  is defined as the derivative of  $F_X(x)$  as

$$f_X(x) = \frac{dF_X(x)}{dx}$$

- There are certain key probability distribution families that have been found very useful in describing the chance structures associated with real life random events
  - Gaussian probability distribution function
    - Bell curve
  - Exponential probability distribution function
    - Time-to-event
  - Binary probability distribution function
    - Heads or tails?
  - Binomial probability distribution function
    - How many heads or tails in so many repeats?
  - Poisson probability distribution function
    - How many heads or tails so far?

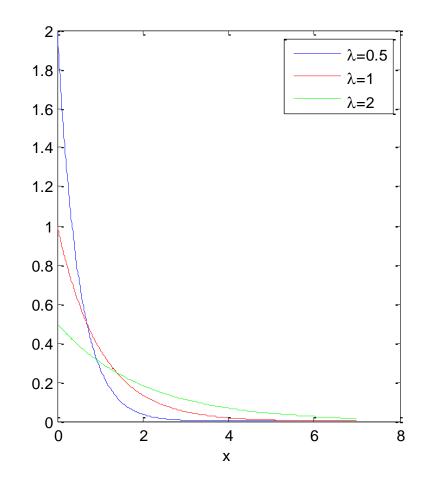
- Gaussian probability
   distribution
  - A continuous function with two parameters
    - Mean  $\mu$
    - Variance  $\sigma^{\rm 2}$

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)}$$



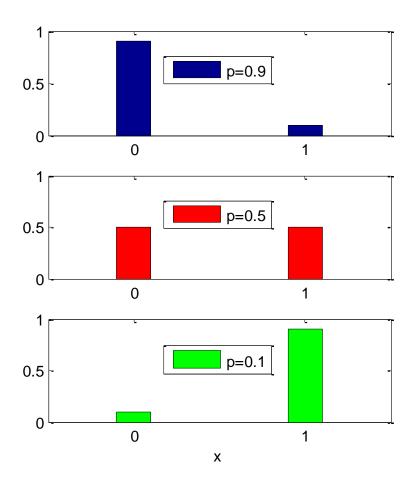
- Exponential probability distribution
  - Another continuous distribution, this time with one parameter
    - The rate of change  $\lambda$
  - This is the only memoryless continuous distribution

 $f_X(x) = \lambda e^{-\lambda x}$ 



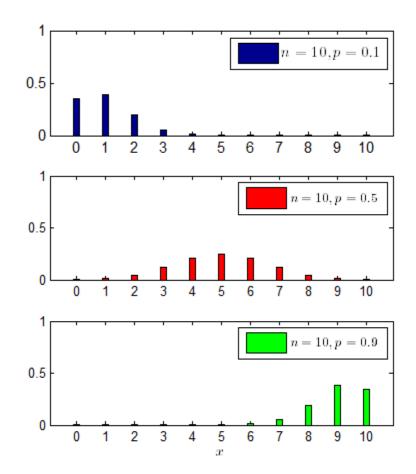
- Binary probability distribution
  - A discrete distribution with only two possible outcomes  $p_X("first outcome") = p$  $p_X("second outcome") = 1 - p$
  - The set of outcomes can be varied
    - {0,1}
    - {-1,1}
    - $\{A, B\}$

• ...



- Binomial probability
   distribution
  - A discrete distribution counting two possible outcomes in so many independent repeats with  $p_X("first outcome") = p$  $p_X("second outcome") = 1 - p$
  - The probabilities are then given by
  - Pr{"k first outcome in n repeats"}

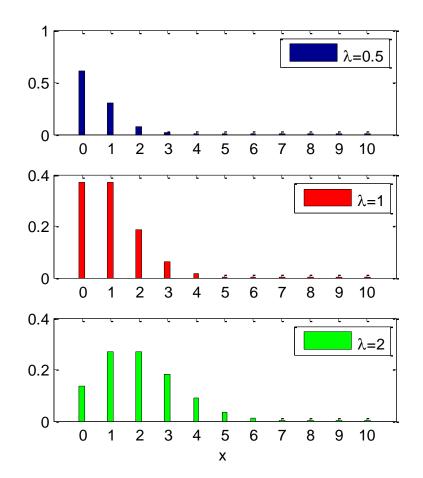
$$= \binom{n}{k} p^k (1-p)^{n-k}$$



- Poisson probability distribution
  - Another discrete distribution with one parameter
    - Rate of change  $\lambda$
  - Counts the number of times an event of interest occurs in a fixed period of time

$$p_X(x) = \frac{\lambda^x e^{-\lambda}}{x!}$$

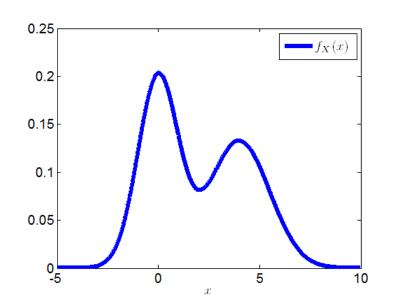
 Interestingly, the time separation between successive events is exponentially distributed



- A sample set represents a collection  $\{x_j\}$ , i = 1, 2, ..., N of values observed from a given random variable
  - A collection of freshman heights from a randomly selected group of 10 first year students
  - Sequence lengths of 10000 randomly selected human proteins
  - Ages (in years) of 120 Alzheimer's Disease patients
  - ...
- The distribution of values in the sample set can be characterized using
  - Histograms
    - $N_k$  represents the number of samples in an interval  $(x'_{k-1}, x'_k]$  with

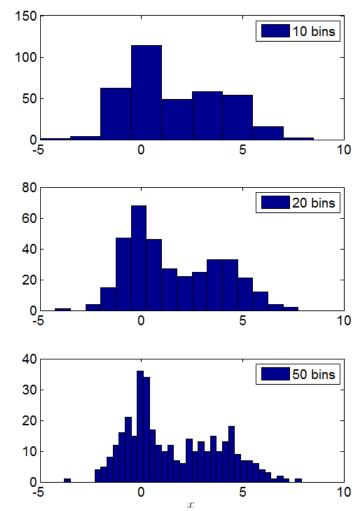
$$x'_0 < x'_1 < \dots < x'_{K-1} < x'_K$$

- *K* represents the number of bins
- Sample statistics
  - Sample mean  $m = \frac{1}{N} \sum_{i} x_{i}$
  - Sample variance  $s^2 = \frac{1}{N-1}\sum_i (x_i m)^2$



**Above:** The probability density function of some random variable *X* 

**Right:** Histograms of 1000 realizations of *X* with different bin sizes



- Remarks:
  - Histograms are informative only when the bins are located and sized appropriately
    - There is no sense in placing the bins on regions of zero occurrence
    - If the bins are too small, the resolution will be high, but they will cover only a few samples producing large errors
    - Larger bins will possess many samples providing a smaller error, but the resolution will be poor
  - Sample mean m and variance  $s^2$  (standard deviation s too) describe where the samples are centered and how wide they are dispersed
    - This is usually fine for unimodal distributions with a single peak
    - On the other hand, this is terribly inadequate to represent multimodal distributions
      - The samples may be clustered around a handful of values with little or no dispersion
      - The mean will not capture this localization, and the standard deviation will indicate large dispersion when there is only very little

# Hypothesis Testing

• Suppose we are given two sample sets

$$\{x_i\}, i = 1, 2, ..., N_x$$

and

$$\{y_j\}, j = 1, 2, \dots, N_y$$

- The heights of freshman students in EE and MB&G
- The sequence lengths of human and yeast proteins
- ...
- The task is to decide if these two sample sets represent events with different characteristics
  - These sample sets represent events with different characteristics if and only if the underlying
    probability distributions are different
- One option it to generate histograms for the two sets and see if they look different
  - Feasible first-attempt, but difficult to infer a statistical significance measure
    - Requires a measure of distance between histograms and permutation tests
- Another option is to assume these sample sets originate from distributions of the Gaussian family with potentially different parameters, and test to see if their parameters might be the same

# Hypothesis Testing

- Presumptions about the statistical nature of the observed data are tested against empirical evidence presented by the data
- Formally:
  - A null hypothesis  $H_0$  is formulated postulating a statement
    - the uninteresting explanation for the observed data
  - A **complementary hypothesis**  $H_c$  is automatically formulated postulating the invalidity of the statement
    - the interesting/desired/hoped-for explanation
  - A probability *P* is computed as the probability of observing the actual observed sample statistic under the null hypothesis
  - If the probability is smaller than a prescribed significance threshold, the null hypothesis is rejected at the benefit of the complementary hypothesis
    - Small *P* values indicate that the sample statistic is unlikely to be observed if null hypothesis were true
    - Typical *P* value thresholds are 5% or 0.1%
- Note that this strategy requires a statistic to be computed from the data with a known distribution under the null hypothesis
  - Any statistic can be used as long as its distribution can be *guessed* well

- Consider the following problem:
  - Two sample sets  $\{x_i\}$  and  $\{y_j\}$  are provided representing the value observed for a parameter of interest from two different groups
    - $\{x_i\}$  are the realizations of a random variable *X*
    - $\{y_j\}$  are the realizations of a random variable *Y*
  - Let  $\mu_X$  and  $\mu_Y$  represent the unknown means of the random variables X and Y
  - The task is to test the validity of the null hypothesis

$$H_0: \mu_X = \mu_Y$$

with a significance threshold  $\alpha \ll 1$ 

 $\rightarrow$  two-sided two-sample *t*-test

- A *t*-test is a statistical comparison test that computes a probability for the null hypothesis given the data
- If the probability is smaller than the prescribed significance α, the null hypothesis is rejected in favor of the complementary hypothesis
- A few variants exist for the *t*-test
  - Equal sample sizes, equal variances
  - Unequal sample sizes, equal variances
  - Unequal sample sizes, unequal variances
  - Paired vs. unpaired
- The test calculates a *T* statistic for each case, and computes its probability when the null hypothesis is true as the *P* value

• For unequal sample sizes, equal variances:

$$T = \frac{m_X - m_Y}{s\sqrt{\frac{1}{N_X} + \frac{1}{N_Y}}}$$

where

$$m_{X} = \frac{1}{N_{X}} \sum_{i} x_{i}, m_{Y} = \frac{1}{N_{Y}} \sum_{j} y_{j}$$

$$s_{X}^{2} = \frac{1}{N_{X} - 1} \sum_{i} (x_{i} - m_{X})^{2}$$

$$s_{Y}^{2} = \frac{1}{N_{Y} - 1} \sum_{j} (y_{j} - m_{Y})^{2}$$

$$s = \sqrt{\frac{(N_{X} - 1)s_{X}^{2} + (N_{Y} - 1)s_{Y}^{2}}{N_{X} + N_{Y} - 2}}$$

and

 $DF = N_X + N_Y - 2$ 

- Procedure for testing for the equality of means:
  - 1. Given the sample sets  $\{x_i\}$  and  $\{y_j\}$
  - 2. Calculate the sample means and variances
  - 3. Calculate the *T* statistic
  - 4. Compare the absolute value of the *T* statistic to the critical value  $T_c$  for which

 $F_t(T_c) = 1 - \alpha/2$ 

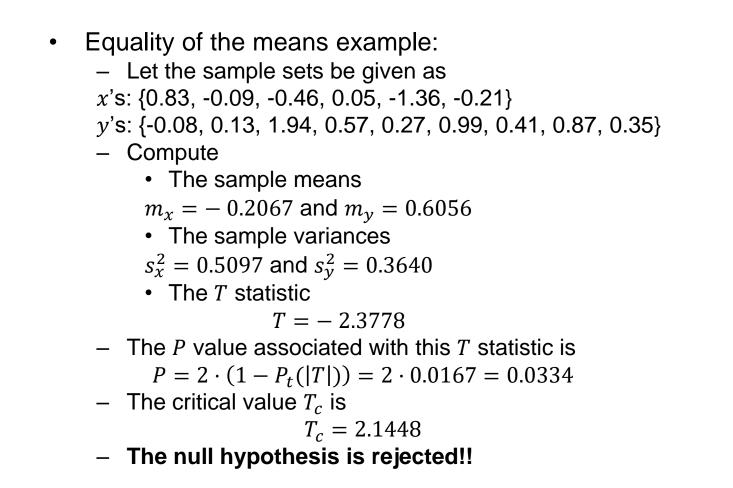
where  $F_t$  denotes the cumulative distribution function of a t random variable with the corresponding degrees of freedom under the null hypothesis

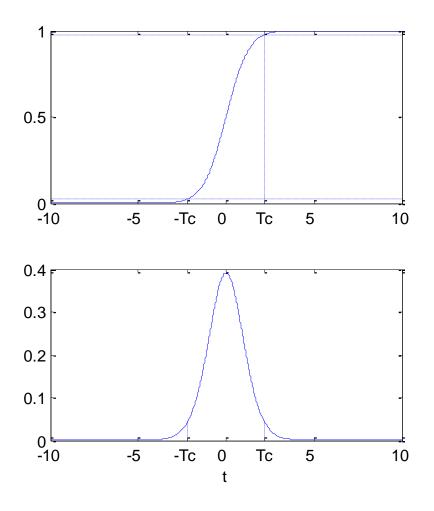
#### OR

Calculate the *P* value via

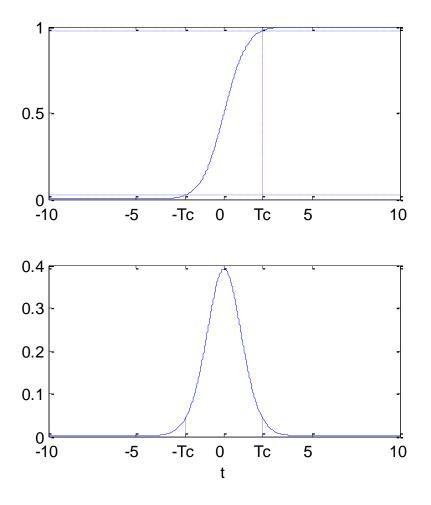
$$P = 2 \cdot \left(1 - F_t(|T|)\right)$$

and see if it is smaller than  $\alpha$ 





Example (continued): ۲ - Now, let the sample sets be given as *x*'s: {0.83, -0.09, -0.46, 0.05, -1.36, -0.21} *y*'s: {-0.08, 0.13, (4.94, 0.57, 0.27, 0.99, 0.41, 0.87, 0.35} Compute \_ The sample means  $m_x = -0.2067$  and  $m_y = 0.9389$  The sample variances  $s_x^2 = 0.5097$  and  $s_x^2 = 2.3648$ • The T statistic T = -1.6914- The *P* value associated with this *T* statistic is  $P = 2(1 - P_t(|T|)) = 2 \cdot 0.0573 = 0.11$ – This time, the null hypothesis is not rejected!! – What is going on??



- Remarks:
  - The *t* test is susceptible to deviations from the presumptions
    - Gaussianity of the underlying distributions
    - Presence of outliers
  - In addition, it determines whether there is reason to believe that the unknown means are different, but says little about how different they are
    - Given sufficient number of samples, the statistical power may suffice to detect even the tiniest differences between the means
    - Conversely, not detecting a difference of the means in a significant manner may simply be because the available data does not provide sufficient statistical power to detect a small difference
  - Finally, it is helpless when the random variables are multivariate
    - Hotelling's  $T^2$  test can be used but is problematic

- Often, several parameters are measured jointly and recorded in experiments
  - Heights, ages, and grade point averages of college freshmen
  - Lengths and amino acid compositions of amino acid sequences of human proteins
  - Gene expression of 40K genes in microarray experiments

- ...

- Such multivariate data sets require multivariate data analysis methods
- A common task when multivariate data from two or more sample sets are present is whether classification rules that distinguish these sets from one another can be constructed
  - If such a rule can be constructed, one can then determine
    - to which group a novel sample should belong
    - which parameter values are critical to distinguish the samples of different groups and in what conditions
  - Both these possibilities are absolutely vital to understand the biological problems in consideration

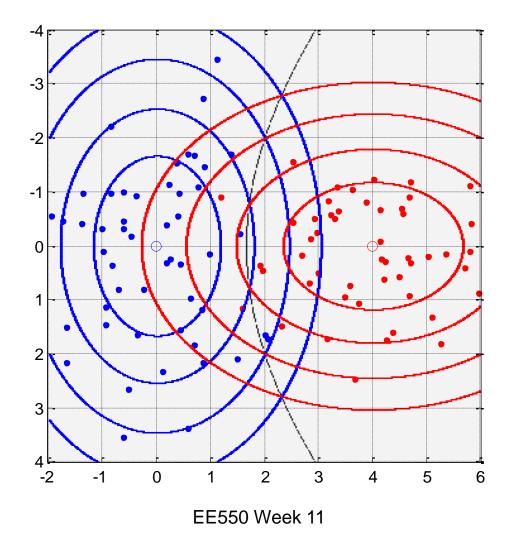
- Given training data, classifier construction strategies are studied under two general categories
  - Parametric classification rules
    - A parametric model is assumed for the underlying multivariate probability distributions of the different groups
    - The parameters for these distribution models are estimated from available data
    - An optimal decision boundary is deduced from the estimated probability distributions
  - Nonparametric classification rules
    - No parameter-based model is assumed
    - Classification rules are constructed based on the similarity and distance structure between the available –manually annotated– "training" samples

- Maximum likelihood classification
  - Given the multivariate training data
  - Estimate the means and the covariance matrices for all sample sets
    - The estimated sample distributions then become multivariate Gaussian distributions with the corresponding mean vectors  $\mu_i$  and the covariance matrices  $\Sigma_i$  as

$$f_i(\mathbf{x}) = \frac{1}{\sqrt{(2\pi)^n |\det(\Sigma_i)|}} e^{-\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma_i^{-1}(\mathbf{x} - \mu_i)}$$

 Construct the classification rule that assigns a new sample to the sample set with the greatest value of the probability density function at the new sample

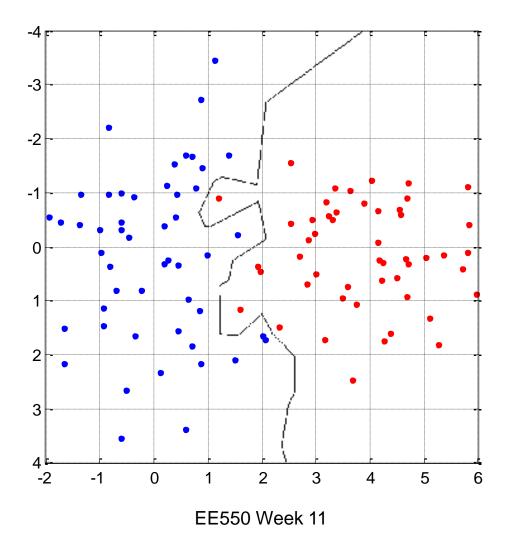
$$f^{\mathrm{ML}}(\boldsymbol{x}) = \arg\max_{j} f_{j}(\boldsymbol{x})$$



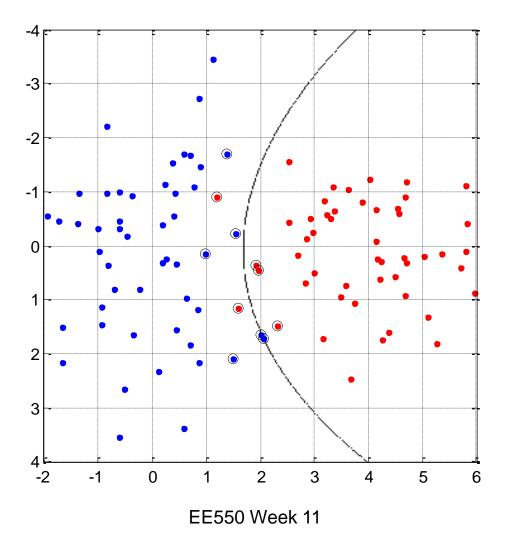
- Nearest neighbor classification:
  - Store all the available data for training in a reference set  $\{x_j, y_j\}, x_j \in IR^n$ ,  $y_j \in \{1,2\}$  with  $j = 1,2, ..., \ell$
  - Assign the newly observed sample to the class with most similar samples
    - Similarity computed in terms of a defined measure, or as inverse distance, or a weighted combination, ...
  - The classification rule is given by

$$f^{\rm NN}(\boldsymbol{x}) = y_{j_0}$$

where  $j_0 = \arg \min_j \rho(x, x_j)$ , with  $\rho(x, x_j)$  calculating the distance between samples x and  $x_j$ 



- Support vector machine classification:
  - A maximum margin linear classifier is constructed to separate the samples of two different classes
  - Nonlinear solutions are obtained by employing an inner product kernel to replace the original inner product between the samples
    - polynomial, Radial Basis Function, sigmoid, ...
  - Linear maximum-margin solution in the transform space corresponds to a nonlinear solution in the observation space
  - For more details, see the literature
    - Maximization of the margin using the method of Lagrange multipliers
    - Karush-Kuhn-Tucker optimality conditions that produce the support vectors
    - Generalization to multiple class problems



# Summary

- Bioinformatics uses statistical analysis techniques to address molecular biology questions emanating from quantitative data in large volumes
  - The data collected from high throughput experiments can only be handled using computational methods
  - These methods use different strategies to answer a variety of questions
    - Whether the nature of measured parameters change from one group to another
    - Whether it is possible to derive classification rules to distinguish the different groups based on the measured data